STRATEGIES FOR SAMPLE ATTRITION ANALYSES IN THE NA-TIONAL SURVEY OF BLACK AMERICANS

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Introduction

Since its inception in 1979-80, the *National Survey of Black Americans* (NSBA) has served as an extensive storehouse of information on African Americans for researchers around the U.S. Studies using the NSBA have focused on numerous topics in a variety of fields, including religion (e.g., Ellison and Sherkat 1990; Sherkat and Ellison 1991), mental and physical health (e.g., Jackson, et al. 1996; Johnson and Broman 1987; Neighbors, et al. 1983), family (e.g., Taylor and Chatters 1991), social support (Lewis 1989; Taylor and Chatters 1986), and health and social services utilization (Neighbors 1985; Neighbors, Musick and Williams 1998; Taylor, Neighbors, and Broman 1989). Clearly, the data set has provided insights into the lives of African Americans that would have been difficult to gain through other mechanisms.

Following the initial wave of data collection in 1979-80, NSBA staff collected data over 3 subsequent waves in 1987-88, 1988-89, and 1992. Researchers have and continue to use these data to examine the multiple topics, including those listed above. However, the NSBA faces a major limitation that is encountered by all longitudinal data sets but is difficult to remedy: sample attrition. By the fourth wave in 1992, approximately 1,456 respondents, or 69% of the original sample size, had departed the sample. Most of the sample attrition (n = 1,301 or 62%) was non-mortality attrition. That is, sample respondents left the survey due to refusal, an inability to participate, or because the survey staff were unable to reach them.

Sample attrition over time is a problem due to the patterned nature of the attrition. For example, a variety of research (e.g., Pappas, et al. 1993, Rogers, et al. 2000, Sorlie, et al. 1995) has shown that mortality over time is non-random; that is, it is patterned by certain factors, such as age, socioeconomic status, and health. Research has indicated that non-mortality attrition is also non-random. Indeed, Wolford and Torres (2001) documented this fact for the NSBA sample. When sample attrition is

non-random and cases from reduced latter waves are used, it is possible that any estimates arrived through regression analyses will be problematic (Breen 1996; Heckman 1979). Consequently, it has been argued that adjustments must be made to the analyses to account for the possible biases that might result. Although, more recently some researchers, notably Stolzenberg and Relles (1997) have argued that techniques used to adjust for attrition bias can be overused and thus produce erroneous results. Nevertheless, it is reasonable to expect that techniques designed to overcome the attrition problem should be used whenever analyzing data from the latter three waves of the NSBA.

The purpose of this paper is to explore the sample selection issue within the NSBA. Drawing on the findings of Wolford and Torres (2001), we construct models to predict selection within the NSBA. However, we go beyond the work of Wolford and Torres by incorporating additional variables that are significant predictors of attrition. Once we establish the attrition models, we provide a substantive example using data from the fourth wave of the NSBA. In this example, we show estimates that are unadjusted for attrition compared to those that are adjusted using two different methods. This example focuses on distress, a topic which has been previously examined in the fourth wave of the NSBA (e.g., Ellison, et al. 1997).

Readers should note that our goal here is to provide a guide to constructing sample attrition models with the NSBA data. Keeping that in mind, we will attempt to provide an explicit rationale for the models and for the variables we include in the models. Moreover, we provide detailed explanations of the variables used in our analyses to facilitate replication and extension in future NSBA studies. Our goal is to show that attrition is a matter of serious concern in the NSBA, but one that can be minimized with the proper procedures.

Predicting Attrition

If we consider that attrition falls into one of two categories, mortality and non-mortality, then the task of searching for predictors of attrition becomes somewhat easier. A number of studies oriented towards understanding the mortality process have documented the factors that often contribute to increased mortality. Here we review several sets of these factors, many of which would overlap with the non-mortality type.

Mortality-based Attrition

Sociodemographics. Generally speaking, research shows that a number of sociodemographic factors are associated with mortality. For instance, men, older adults, those living in urban areas, and unmarried adults all carry greater risk of mortality (Lantz, et al. 1998; Rogers, Hummer and Nam 2000; Smith, et al. 1995). Other

research (e.g., Kaplan 1996) has suggested that household characteristics, such as crowding, might have effects on individual mortality rates as well. Consequently, studies of attrition should attempt to account for these sorts of factors.

Socioeconomic Status. One set of factors receiving a great deal of attention in the mortality process is socioeconomic status (SES). In essence, research using populations both in the U.S. and abroad shows that adults with lower SES tend to have worse health and higher mortality rates (e.g., Lantz, et al. 1998; Rogers, Hummer and Nam 2000). In these studies, SES is measured in a variety of ways, including education level, family or personal income, occupation, and wealth. Given the strong association between SES and mortality noted in previous studies, it is likely that these factors will emerge as significant predictors of attrition within the NSBA.

Employment. Related to the research on socioeconomic status, studies of the effects of unemployment and work factors have shown the importance of these factors for longevity. Findings indicate that adults who are unemployed or who face hazardous or stressful work conditions are at greater risk for mortality compared to those who work and have jobs that are characterized by safe and stable work environments (Smith, et al. 1998). Without firm data on the nature of the work environment, it can be difficult to ascertain the nature of the work involved. However, through measures such as job prestige and similar factors, we can arrive at approximations of the work experience.

Community Involvement and Social Integration. A number of studies have focused on the role that community involvement and social integration play in the mortality process. For example, Moen and her colleagues (1989) showed that women who were active members of an organization had lower mortality rates than their uninvolved counterparts over a 30-year follow-up period. In their work in Tecumseh, Michigan, House and his colleagues (1982) showed similar findings for participation in organizational meetings. Another set of research focuses specifically on involvement in religious institutions and how those affect the mortality process. In general, these studies show that respondents who are more involved in the church tend to live longer than those who are less involved (e.g., Hummer, et al. 1999; Strawbridge, et al. 1997).

Health. Perhaps one of the strongest predictors of mortality shown through numerous studies is respondent health status. Most studies of mortality show that poor health, measured along multiple dimensions, is predictive of greater mortality risk. More specifically, authors have noted strong associations with self-rated health (Greiner, et al. 1996; Idler and Benyamini 1997; Idler and Kasl 1991; Wolinsky and Johnson 1992), functional health (Rogers, Hummer and Nam 2000), and the presence of chronic health conditions (Musick, Herzog and House 1999). Any study purporting to predict sample attrition over time, especially in terms of mortality, should adjust for these factors.

Non-mortality Attrition

Ordinarily, non-mortality forms of attrition break down into three main forms. First, respondents may have departed the sample because they simply refused to cooperate with the additional waves. Second, data staff might be unable to locate, and therefore interview, respondents. Third, respondents might be incapacitated, either through health problems, institutionalization, incarceration, or enlistment in military service rendering them difficult if not impossible to interview. Given these three forms of non-mortality attrition, we should look for sets of factors that would apply to one or more.

Of the factors outlined above for mortality, several non-mortality predictors are possible. First, good health in the first wave might be indicative of people who would not be incapacitated or institutionalized (e.g., in a nursing home) during later waves. Consequently, we would expect that healthy first wave respondents would more frequently appear in latter waves.

Second, high levels of socioeconomic status should lead people to be more cooperative with the survey. Indeed, we know that higher SES is associated with other forms of community participation, such as volunteering and participating in other formal organizations (Wilson and Musick 1997); it is likely the same processes underlying the linkage between SES and these activities would hold true for survey research. In addition, higher levels of SES might be indicative of stability. Given that much of the sample attrition in the NSBA is due to the inability of the survey staff to relocate respondents, those with greater residential stability in the first wave will be more likely to be relocated during subsequent waves. In this regard, whether the respondents own their homes is a potentially strong predictor given that homeowners are more tied to their residences and thus are less likely to move. In fact, Wolford and Torres (2001) found that one of the strongest predictors of attrition in the NSBA was home ownership. Another indicator of SES, education, might have substantial impacts on attrition in the following way: more educated respondents might place a higher value on the survey and its goals leading to a greater willingness to participate.

Third, active involvement in the community should predict greater retention. It should be the case that respondents who are engaged in their community will feel more a part of the community and therefore will have more residential stability and be easier to locate. Further, even if involved respondents had moved, it is possible that friends and neighbors in the community will know enough about respondents' new residences to help relocation efforts. Active involvement in community life might also be indicative of an underlying proclivity to participate in voluntary activities, such as re-

search studies. In this regard, work by Ellison (1992) indicates that more religious African Americans tend to be more cooperative in interview settings. Consequently, we should see fewer refusals among active community members, and especially religious respondents, than among those who are not as active.

Other factors, such as employment status and residence might also play a part in predicting non-mortality attrition. Respondents with steady jobs will be those less likely to relocate residences. Further, given the tight-knit nature of less urban communities (Beggs, Haines and Hurlbert 1996), people living in less urban areas may be located relatively easily through friends and family even if they have changed residences. In sum, there are a number of ways in which the factors listed above could affect sample attrition both in terms of mortality and non-mortality mechanisms.

The Utility of Interviewer Observations

Researchers can use variables in many of the areas listed above in their studies to control the effects of sample attrition. A problem arises, however, when the substantive outcome of the study is predicted by many of the same variables as the sample attrition. For example, given the processes outlined above, a researcher wanting to study changes in health across the waves would probably use many of the variables listed above to predict changes in health status. If those same variables are used in the selection models, a problem could arise. More specifically, the variables used in the selection model must be somewhat different than those in the substantive model, otherwise the selection model will be unable to effectively adjust for the selection model overlap with the variables in the substantive model, the predictive power of the selection portion rests only in the non-linearity in the probit model used to compute the selection adjustments. Consequently, a major goal for constructing proper sample selection models is to use as many non-overlapping factors as possible in the selection adjustantive models.

One source of non-overlapping factors that often go unused in surveys is the interviewer observations. Interviewers are commonly asked to assess the respondents' demeanor during the interview, the state of their home and neighborhood, their health, and their overall apparent quality of life. Researchers can tap these measures as possible predictors of selection as many of the variables fall into the same categories as those listed above. In general, few researchers have used these measures in the NSBA. Yet, in the few cases where these measures have been used, they have yielded valuable information (e.g., Ellison 1993; Hughes and Hertel 1990; Keith and Herring 1991). On the topic of attrition, Wolford and Torres (2001) used interviewers' observations of respondents' income in the first wave of the NSBA. Such a measure could be used as an independent appraisal of the respondents' overall SES and can be used

in conjunction with the SES measures provided by respondents themselves. Other measures in the NSBA ask interviewers to rate how interested respondents seemed in the interview, whether they were impatient, were having difficulty hearing words, or were reluctant about signing the re-contact form. In short, variables such as these potentially contain great utility for predicting attrition. Moreover, because they are unlikely to be used in predicting most substantive topics, their usage will not overlap with the variables in the main models.

Data and Methods

Sample

The data used in this analysis are taken from all four waves of the *National Survey of Black Americans*, a probability sample of 2,107 African Americans (Jackson and Gurin 1997). The first wave of the study was conducted using face-to-face interviews in 1979-80. Follow-up surveys were conducted using phone interviews in 1987-88 (N=931), 1988-89 (N=785), and 1992 (N=651).

Attrition Measures

Socioeconomic Status. Socioeconomic status was measured using four variables. The first measure (<u>Home Ownership</u>) is coded one if respondents own their homes and zero otherwise. <u>Education</u> is categorized into four groups: (1) 0 to 11 years, (2) 12 years or high school graduate, (3) 13 to 15 years or some college, and (4) 16 years or more or college graduate. <u>Personal Income</u> is grouped into eight categories ranging from 0 to \$15,000 or more. Interviewers' estimation of respondent income (<u>Low Income – IW</u>) was coded one if interviewers felt the respondent's income fell below \$10,000 per year and zero otherwise. The final measure (<u>No Problems Paying Bills</u>) asks how much respondents worry that their total family income will not be enough to meet expenses and bills. The response categories range from (1) a great deal to (4) not at all.

Sociodemographics. All models were adjusted for the effects of sociodemographic characteristics including gender (<u>Female</u>: 0 = male, 1 = female), living in a three generation family (<u>Three Generation Family</u>: 0 = no, 1 = yes), number of persons living in the household (<u>Number in HH</u>) ranging from 0 to 13, and living in a self-representing urban area (<u>Urban Residence</u>: 0 = no, 1 = yes).

Employment. The first measure of employment (<u>Employment Search Status</u>) is coded one if the respondent was laid off or not working at all and zero if otherwise. The second measure (<u>Employed only Part of 1978</u>) is coded one if the respondent did not work for some weeks in 1978 and zero if otherwise. The most important factor in a

job (<u>Most Important Job Factor</u>) is coded one if the respondent reported friendly people to work with and zero if otherwise. Frequency of absence from work (<u>Absence from Work</u>) is coded one if the respondent reported missing work very often and zero otherwise. Finally, wage type (<u>Paid a Salary</u>) is coded one if the respondent is paid a salary and zero otherwise.

Community Involvement. Level of community involvement is measured using three variables. Whether or not the respondent voted in the last presidential election (<u>Voted</u> in <u>Presidential Election</u>) is coded one for yes and zero if otherwise. Frequency of church activity besides attendance at regular services (<u>Church Activity</u>) is coded one for nearly everyday or more zero otherwise. An additional variable is included to measure whether or not the topic of the study affected nonresponse. Whether or not the respondent would vote for a candidate with the best platform for blacks even if he/ she did not belong to the respondent's party (<u>Supports Black Platform</u>) is assigned a value of one for yes and zero otherwise.

Interviewer Observations. Several measures were used in order to assess individual's affective states and attitudes toward the survey (as observed by the interviewer) at baseline on the likelihood of their participation in subsequent waves. These include whether or not the respondent ever asked how much longer the interview would take (Asked Interview Length: 0 = no, 1 = yes), ever had difficulty with any of the wording used in the interview (Difficulty with Words: 0 = no, 1 = yes), was reluctant to sign the recontact sheet (Reluctant to Sign Re-contact: 0 = no, 1 = yes), and if the respondent's home appeared to be in need of minor (Minor Repairs Needed: 0 = no, 1 = yes) or major repairs (Major Repairs Needed: 0 = no, 1 = yes). For the latter two variables, the reference category was no repairs appeared needed. A three-item index of physical problems (Physical Problems) was coded one if the respondent had hearing problems, vision problems such as blindness or unusually thick lenses, or physical impairments such as missing limbs, artificial limbs, facial scars, etc. and zero if no problems were reported.

Psychological Distress Example

The variables gender, employment search status, education, and problems paying bills as described above are also included in the distress example. Additional measures included in the distress example are grouped into the following categories: socio-demographics, religion, social integration, and health and well-being.

Psychological Distress. Distress was measured in the fourth wave using an index created by taking the arithmetic mean of ten items: (a) under strain, stress or pressure; (b) in low or very low spirits; (c) been moody or brooded about things; (d) felt downhearted and blue; (e) feel depressed; (f) felt tense or high-strung; (g) able to relax

(reverse-coded); (h) bothered by nervousness or your nerves; (i) felt restless and upset; and (j) been anxious and worried. For each of the items, respondents were asked how much they had felt that way over the past month. Responses range from (1) none of the time to (4) all of the time.

Sociodemographics. All models were adjusted for the effects of certain social and demographic characteristics including, age (Age: in years), marital status (Married: 0 =not married, 1 = currently married), and family income (Family Income: in categories ranging from 0 to \$30,000 or more).

Religion. We include three measures of religious attendance and practice. First, <u>Religious Attendance</u> asks respondents how often they usually attend religious services. The five response categories include (1) less than once a year or never, (2) a few times a year, (3) a few times a month (1 to 3 times), (4) at least once a week (1 to 3 times), and (5) nearly everyday (4 or more times a week). The second and third variables indicate how often the respondent prays. The responses range from (1) never to (5) nearly everyday or 4 or more times per week. It should be noted that we centered the prayer variable (<u>Prayer</u>) and then squared it (<u>Prayer</u>²) to estimate curvilinear effects.

Social Integration. The first measure of social integration (<u>Confidants</u>) indicates how many friends, not including relatives, the respondent feels free to talk with about their problems. The response categories range from (1) none to (4) many. The second measure (<u>Social Interaction</u>) asks how often the respondent sees, writes, or talks on the telephone with friends. The response categories range from (1) hardly ever or never to (6) nearly everyday or 4 or more times per week. Finally, we include a two-item index of <u>family support</u> that asks respondents (a) how often people in their family help them out, and (b) how much help they are. The response categories range from (1) never to (4) very often, and from (1) only a little help to (4) a great deal of help, respectively.

Health and Well-Being. Health and well-being are measured using three variables. In order to measure the degree of personal problems, respondents were asked whether or not they had ever (a) had a personal problem they could not handle themselves, (b) felt down and depressed, so low they could not get along, (c) had a personal problem where they felt so nervous they could not do much of anything, or (d) felt they were about at the point of a nervous breakdown. We created a new variable (<u>Mental Health Problems</u>) with response categories that range from (1) never experienced any of these personal problems to (5) felt they were about at the point of a nervous breakdown (categories a, b, and c from above were coded 2, 3, and 4, respectively). <u>Self-esteem</u> is a six-item index created using the following items: (a) I am a useful person to have around; (b) I feel that I'm a person of worth; (c) As a person, I do a good job

these days; (d) I feel I do not have much to be proud of; (e) I feel that my life is not very useful; and (f) I feel that I can't do anything right. For each item, respondents were asked to indicate whether the statement was (1) never true, (2) not often true, (3) often true, or (4) almost always true of them. The latter three items were reverse coded so that higher scores are indicative of greater self-esteem. Finally, functional health at the first wave (Wave 1 Functional Health) is a sum of the number of the following debilitating conditions that keep the respondent from working or carrying out their daily tasks: arthritis/rheumatism, ulcers, cancer, high blood pressure, diabetes, liver problems, kidney problems, stroke, circulation problems, sickle cell anemia, or any other health problems.

Research Plan

In the remainder of this paper we examine the attrition process in the NSBA using data from the first wave. Given that Wolford and Torres (2001) have already done a substantial amount of work on the topic, we use their paper as a starting point in our own work. Our plan of analysis is as follows. First, we show that the sample attrition in the NSBA is indeed non-random and thus should be taken into account when using data beyond the first wave. Second, we attempt to replicate the Wolford and Torres (2001) attrition prediction models for both all-cause and non-mortality attrition using logistic regression. Although we were unable to completely replicate their findings, we have made efforts to come as close as possible to doing so. After presenting the Wolford and Torres models, we make adjustments by removing several variables that are not predictive of attrition and add others, especially interviewer observations that we found to be predictive of attrition across multiple waves. We further show the improvements in model fit and predictive ability that are made based on these new model specifications.

Finally, we examine a model that predicts distress in the fourth wave based on covariates in the first. In this section, we estimate a model that is unadjusted for attrition, one that is adjusted using the Heckman two-step least squares estimator, and a third using the Heckman two-step maximum likelihood estimator. We then compare coefficients across models and note the changes made through different forms of adjustment. Note that for the recoding of data and logistic regression models, SAS 6.12 for Windows was used. Because SAS is not ideal for computing attrition models, we rely on LIMDEP 7.0 for that portion of the analyses. For readers without access to LIMDEP but who can use SAS, a macro program is available from the SAS website (www.sas.com) that can be used to estimate the models.

Results

NSBA Attrition: Non-Random?

In Wolford and Torres's (2001) paper, the authors showed a few examples of the ways in which the NSBA sample is affected through attrition. In Table 1, we provide more details on this issue. More specifically, we report mean levels of the variables used in the attrition analyses across all four waves of the sample. We would expect that if the attrition across the waves was completely random, the means for each of the variables would remain virtually unchanged. Note that the means reported for each wave are the means of the first wave variables for those remaining in each wave.

As one can see looking across the table, the means of the variables do indeed change as the sample becomes more reduced in size. For example, in terms of sociodemographic factors, the sample becomes more female and rural over time. Regarding socioeconomic status, the sample rises appreciably in terms of education and income. Indeed, the proportion of the sample owning homes between waves one and four rises by fifteen percentage points. We also note large changes in employment status across the four waves. Likewise, some changes in community involvement are apparent, such as voting rates increase as does church activity and support of a Black platform. In short, across all sets of factors used to predict attrition, we observe changes in mean levels indicating fundamental changes in the composition of the sample.

The implications of the change in sample composition are twofold. First, it is readily apparent that by the fourth wave, and indeed even the second, the sample is no longer representative of the population from which it was drawn. Unfortunately, this problem cannot be resolved using sample selection methods; however, other techniques, such as weighting, could reduce the problem of representativeness. Second, and more importantly for our purposes, the changes indicate that the attrition has been non-random. That is, it has been patterned in a variety of ways. Unless we take account of these selection patterns in our analyses of the latter waves of data, biased and inconsistent estimates are possible (Breen 1996).

Predictors of All-cause Attrition

Most researchers using the NSBA data will be concerned with all-cause attrition, that is, attrition from both mortality and non-mortality sources. Because the attrition models must take account of all attrition to function correctly, these are the most appropriate outcomes to model. We show the results of these models in Table 2. All of the variables in the sociodemographics, socioeconomic status, employment, and community involvement sections were those used in Wolford and Torres's (2001) final

	Range	Mean Wave 1 R's (n = 2,107)	Mean Wave 2 R's (n = 931)	Mean Wave 3 R's (n = 785)	Mean Wave 4 R's (n = 651)
% Retained Sociodemographics	_	_	44.19%	37.26%	30.90%
Female	0 - 1	.62	.66	.66	.68
Three generation family	0 - 1	.53	.56	.55	.57
Number in HH	1 – 13	3.18	3.24	3.22	3.27
Urban residence	0 – 1	.47	.44	.42	.44
Socioeconomic Status					
Education	1 - 4	1.90	2.09	2.11	2.18
Personal Income	1 - 8	4.15	4.47	4.53	4.73
Low Income – IW	0 - 1	.45	.35	.34	.31
Home Ownership	0 - 1	.48	.58	.62	.63
Problems paying bills	1 - 4	2.94	3.03	3.03	3.08

Table 1. Ranges and Means of Variables by Wave.

Table 1 (continued). Ranges and Means of Variables by Wave.						
	Range	Mean Wave 1 R's (n = 2,107)	Mean Wave 2 R's (n = 931)	Mean Wave 3 R's (n = 785)	Mean Wave 4 R's (n = 651)	
Employment						
Employment search status	0 – 1	.42	.31	.30	.27	
Employed only part of 1978	0 – 1	.56	.46	.45	.41	
Most important job factor	0 – 1	.07	.06	.06	.06	
Absence from work	0 – 1	.02	.01	.01	.01	
Paid a salary	0 – 1	.00	.01	.01	.01	
Community Involvement						
Voted in presidential election	0 – 1	.54	.63	.64	.66	
Church Activity	0 – 1	.18	.21	.23	.23	
Supports Black Platform	0 – 1	.62	.68	.69	.71	
Interviewer Observations						
Asked interview length – IW	0 – 1	.22	.18	.16	.15	
Difficulty with words – IW	0 – 1	.14	.11	.09	.08	
Reluctant to sign re-contact – IW	0 – 1	.04	.02	.02	.02	
Major repairs needed – IW	0 – 1	.17	.11	.10	.10	
Minor repairs needed – IW	0 – 1	.35	.30	.29	.27	
Physical problems – IW	0 – 1	.10	.05	.05	.05	

	Wave 2 In Sample ^b		Wave 3 In Sample		Wave 4 In Sample	
	Model 1	Model 2	Model 1	Model 2	Model 1	Model 2
Sociodemographics						
Female	1.64***	1.70***	1.63***	1.70***	1.93***	1.93***
Three generation family	1.27*	1.29**	1.17	1.18 +	1.35**	1.35**
Number in HH	1.02		1.01		1.03	
Urban residence	.71***	.72***	.65***	.67***	.74**	.78*
Socioeconomic Status						
Education	1.26***	1.21***	1.23***	1.18**	1.27***	1.23***
Personal Income	1.00		1.00		1.04	
Low Income – IW	.72**	.87	.70**	.86	.69**	.82**
Home Ownership	1.53***	1.50***	1.73***	1.70***	1.66***	1.65***
No Problems paying bills		1.09+		1.06		1.15**
Employment						
Employment search status	.56***	.53***	.54***	.52***	.61**	.48***
Employed only part of 1978	.84		.86		.71*	
Most important job factor	.70+	.69+	.59**	.57**	.69+	.69+
Absence from work	.45*	.43*	.50+	.47*	.39*	.34*
Paid a salary	4.92		6.75+		2.29	

Table 2. Estimated Net Effects of First Wave Covariates on Sample Retention vs. All-Cause Attrition (Logistic Regression Estimates; n = 2,107).^a

models, except for problems with bills in the socioeconomic section. We add that variable as well as the interviewer observations near the bottom of the table to the models. In the first two columns, we regress retention (i.e., respondents coded one if they were in the sample at the second wave and zero otherwise) on the Wolford and Torres variables. According to this model, we find a number of significant predictors of retention, notably home ownership, urban residence, and employment search status. Note that the estimates shown are odds ratios, meaning that values over one are indicative of greater retention based on higher levels of the predictor variable. For example, respondents who are home owners in the first wave are about 53% more likely to be in the second wave sample than those who do not own homes in the first wave. Similarly, women are more likely to be retained than men, as are those who are more educated and who voted in the presidential election. In contrast, people living in urban areas and those with low interviewer ratings of income were less likely to be retained. Looking at the fit statistics of this model, we note that the model explains about 12% of the variance in the dependent variable and is able to accurately predict cases about 69.9% of the time.

In the second column we present a similar model to the first but remove several variables and add others, notably the interviewer observations. The variables we chose for exclusion in the second model were those that were not significant predictors of retention across more than one of the outcomes in either the all-cause or non-mortality attrition cases. These variables include number in household, personal income, employed only part of 1978, paid a salary, and supports Black platform. Even though we add several variables, with these subtractions the degrees of freedom used by the model only increases by two. Looking at the estimates in the second column we note that all except church activity and difficulty with words are significant predictors of retention. The estimates of the Wolford and Torres variables do not change substantially, although in some cases, such as gender and voting, they do change somewhat. Among the new variables, respondents who had no problems paying bills were more likely to be retained. In contrast, those who asked about the interview length, were reluctant to sign the re-contact form, who had physical problems, or who lived in premises in need of repair were all less likely to be retained in the second wave.

Although these new variables are significant predictors themselves, it is important to determine whether they actually improve the overall fit and predictive power of the model. Looking at the fit statistics in column two, we see some improvement is made. The change in chi-square ($D X^2 = 49.67$) is significant, indicating better model fit. The change in R² reflects that the model explains slightly more variance. Further, the new model is better able to accurately predict retention by about 1.8%. In sum then, the second model is an improvement over the first and entails little loss in parsimony.

Voted in presidential election	1.42***	1.46***	1.40**	1.46***	1.35**	1.47***
Church Activity	1.25 +	1.19	1.33*	1.27 +	1.36*	1.25 +
Supports Black Platform	1.18		1.18		1.27*	
Interviewer Observations						
Asked interview length – IW		.75*		.65***		.56***
Difficulty with words - IW		.86		.73+		.65*
Reluctant to sign re-contact -	IW	.40***		.30***		.49*
Major repairs needed – IW		.58***		.56***		.62**
Minor repairs needed – IW		.67***		.65***		.62***
Physical problems – IW		.56**		.61*		.56**
INTERCEPT	-1.02	85	-1.28	97	-2.08	-1.76
$X^2 / d.f.$	265.85/16	315.52/18	282.66/16	341.85/18	305.68/16	358.34/18
$\supset X^2$		49.67***		59.19***		52.66***
R^2	.12	.14	.13	.15	.14	.16
$ ightarrow \mathbf{R}^2$.02		.02		.02
% Predicted Correctly	69.9	71.7	71.0	73.1	72.7	74.6

Notes: ^a Dependent variable is scored one if respondents remained in the sample during the wave indicated and zero if they had dropped out for *any* reason.

^b Odds ratios are shown.

+ p < .10; * p < .05; ** p < .01; *** p < .001

The second and third sets of models follow the same patterns as those shown for the first. That is, both the Wolford and Torres and new variables all predict retention in the third and fourth waves in fairly the same fashion as in the second wave. However, the changes in the models made in the third and fourth waves incur somewhat greater fit and precision than that in the second wave. That is, in the third wave, the new model incurs a change of 59.19 in X^2 and about 2.1% in accurate prediction. The numbers are similar for the fourth wave (D $X^2 = 52.66$; prediction = +1.9%). In sum, for all-cause attrition, the changes we have made to the models do provide better fit and predictive ability. Further, given that the interviewer variables are unlikely to overlap with variables used in a substantive model, the new model could be applied in a number of situations to predict attrition.¹

Non-mortality Attrition

In Table 3 we display the results of the non-mortality attrition models. The estimated effects shown in these models are very similar to those shown for the all-cause attrition models. This similarity should not be surprising given that a large majority of the sample attrition is due to non-mortality causes. Like the all-cause mortality models, the most powerful predictors of non-mortality attrition are SES factors such as education and home ownership and some of the interviewer observations. Likewise, the model fit statistics and changes between the Wolford and Torres models and our own are similar. Consequently, in the interest of conserving space, we do not further discuss the pattern of these findings.

Distress Example

In Table 4 we display the results of our distress example. In the first column, we report the effects of the first wave covariates on fourth wave distress modeled using ordinary least squares regression. Note that our choice of variables in this example was theoretically derived only to the extent that we wanted to find a set of variables that would have some effect on the fourth wave outcome. Consequently, we chose variables along several dimensions known to influence distress, such as gender, so-cioeconomic status, prior mental health problems, current physical health problems, self-esteem, and religious activity. Very few of these variables overlap with the vari-

¹We considered the possibility that self-reported health problems would predict attrition given that some of the attrition is likely due to mortality or an incapacity to participate. Consequently, we examined whether the health measure used in the distress example also predicted attrition. We found that in both the all-cause and non-mortality cause attrition models, this index had no significant effect. This non-significant finding for self-reported health underscores the utility of the interviewer observations, especially for health problems.

(Logistic Regression Estimates). ^a	(n = 2,005) Wave 2 In Sample ^b		(n = 1,994) Wave 3 In Sample		(n = 1,952) Wave 4 In Sample	
	Model 1	Model 2	Model 1	Model 2	Model 1	Model 2
Sociodemographics						
Female	1.55***	1.62***	1.55***	1.62***	1.78***	1.79***
Three generation family	1.32**	1.33**	1.22+	1.23*	1.39**	1.39**
Number in HH	1.00		1.00		1.01	
Urban residence	.70***	.70***	.65***	.65***	.74**	.78*
Socioeconomic Status						
Education	1.26***	1.21***	1.24***	1.19**	1.26***	1.23***
Personal Income	.99		.99		1.04	
Income – IW	.72**	.89	.71**	.88	.70**	.84
Home Ownership	1.66***	1.63***	1.88***	1.83***	1.84***	1.82***
No Problems paying bills		1.11*		1.08		1.16**
Employment						
Employment search status	.60***	.59***	.58***	.57***	.66*	.54***
Employed only part of 1978	.86		.88		.73*	
Most important job factor	.73+	.71+	.61*	.59*	.71	.71
Absence from work	.43*	.41*	.49+	.47*	.36*	.33*
Paid a salary	4.51		6.21+		2.10	

 Table 3.
 Estimated Net Effects of First Wave Covariates on Sample Retention vs. Non-Mortality Attrition

Community Involvement

Voted in presidential election	1.48***	1.54***	1.46***	1.53***	1.42**	1.56***
Church Activity	1.31*	1.25 +	1.40**	1.33*	1.46**	1.34*
Supports Black Platform	1.20 +		1.20 +		1.30*	
Interviewer Observations						
Asked interview length – IW		.74*		.65***		.57***
Difficulty with words - IW		.87		.74+		.64*
Reluctant to sign re-contact -	IW	.38***		.29***		.45*
Major repairs needed – IW		.57***		.55***		.61**
Minor repairs needed – IW		.68***		.67***		.62***
Physical problems – IW		.61**		.68+		.61*
INTERCEPT	95	92	-1.21	-1.06	-1.99	-1.78
X ² / d.f.	257.96/16	307.32/18	276.69/16	335.36/18	293.23/16	347.20/18
D X ²		49.36***		58.67***		53.97***
R ²	.12	.14	.13	.15	.14	.16
D R ²		.02		.02		.02
% Predicted Correctly	69.9	71.7	71.1	73.2	72.6	74.6

Notes: ^a Dependent variable is scored one if respondents remained in the sample during the wave indicated and zero if they had dropped out for *non-mortality* reasons.

^b Odds ratios are shown.

+ p < .10; * p < .05; ** p < .01; *** p < .001

	OLS	Sample	Sample
	Regression	Selection	Selection
		LS Estimates	ML Estimates
Sociodemographics			
Female	.07	.09+	.08+
Age	01***	01***	01***
Married	.04	.04	.04
Education	.00	.01	.01
Family Income	01	00	00
Unemployed	.06	.04	.05
No Problems Paying Bills	06**	05*	05**
Religion			
Religious Attendance	01	01	01
Prayer	08	08+	08+
Prayer ²	04*	04*	04+
Social Integration			
Confidants	.02	.02	.02
Social Interaction	.01	.01	.01
Family Support	.01	.01	.01
Health and Well-Being			
Mental Health Problems	.05***	.05***	.05***
Self-Esteem	03***	03***	03***
Wave 1 Functional Health	.52***	.52***	.52***
Intercept/Constant	2.09	1.94	1.97
Adjusted R ²	.18	.18	

Table 4. Estimated Net Effects of Selected Wave 1 Variables on Wave 4 Psychological Distress.

Notes: * p <.05; ** p <.01; *** p <.001

ables in the attrition models; consequently, we can be certain that our attrition models are having the intended effect.

In Table 4 we display the results of the distress models. In the first model we report the estimates from an uncorrected ordinary least squares regression model. In the second column we show the estimates generated from a Heckman two-step least squares sample selection model, while in the third the estimates are derived from a sample selection model with maximum likelihood estimation. According to Breen (1996) the latter of these models tends to provide the best estimates of the population parameters.

In comparing coefficients across the models, we note one important pattern: very little change is evidenced through the incorporation of the sample selection analyses. Indeed, only three effect sizes, those for gender, income, and employment status, changed between the first and latter two models. Although the effect of gender became significant (p < .05) in the sample selection models, the absolute change in coefficient size was small (D b = .02). Another variable, the squared form of prayer, dropped slightly in significance. However, overall, incorporating the selection models had very little effect on the analyses.

Discussion

Our analyses revealed three important findings. First, the attrition in the NSBA over the four waves is nonrandom; consequently, analyses using latter waves of the sample may be biased unless the nonrandom attrition is accounted for. It should be noted that the approximate 8 to 9 year gap between the first and second waves of the NSBA largely accounts for the exceptionally high level of attrition in wave 2. This long gap exacerbates the already difficult process of relocating respondents that is experienced by all longitudinal surveys. It suggests that new data collection efforts that anticipate any form of follow-up interviewing should plan to recontact respondents over shorter periods of time. Second, the inclusion of interviewer observations is important for the prediction of attrition. Further, because these variables will not often be used in substantive analyses, they will not face the problem of overlap with other variables that are often used. Our findings also suggest that interviewers are fairly skilled at recognizing potential problems with certain respondents. Given declining participation and cooperation rates in all surveys including those with face-to-face interviews, these analyses suggest the need for additional research regarding the dynamics of interview processes. Finally, even though the attrition is non-random, once it is taken into consideration, the coefficient estimates in our example change very little.

Generally speaking, we found several groups of factors to be predictive of attrition in this sample. One of the strongest sets of predictors involved socioeconomic status

measured in several ways. For instance, we found across both forms of attrition and all waves that education, and especially home ownership, were strong predictors of attrition. These findings support our assertion that those SES factors promote stability and perhaps a willingness to advance the goals of the study. Several sociodemographic variables were also predictive, notably gender and urban residence. With regard to gender, it is somewhat unclear why women were less likely to drop out of the sample. In order to better understand this finding, we examined the actual reasons for attrition during the second wave, which had the largest amount of attrition. In reviewing these reasons by gender, we noted that 12.6% of men but only 8.9% of women were not recontacted due to the "too ill/circumstantial" reason category. Based on this difference, it may be the case that men were sicker than women at later waves. Likewise, other analyses show that men were more likely to have died (men: 10.9%, women: 6.0%). However, it might also be the case that men were prevented from participating more often than women due to incarceration. Data from the Department of Justice (Snell 1995) supports this assertion showing that the incarceration rate as of 1993 for Black men was 4.63% compared to only .24% for Black women.² In short, it appears that the differences in attrition we observe are either due to health reasons, death, or some other incapacitation, such as incarceration.

The final set of factors showing great utility for predicting attrition are the interviewer observations. Several of these factors, notably home repair status, physical problems, and reluctance to sign the re-contact sheet were strong predictors of attrition. Indeed, the addition of these factors significantly improved the overall fit and predictive ability of the model. These findings buttress the few studies that have profitably used these measures in the past. They further suggest that future studies on health and other outcomes should consider their usage both in terms of modeling substantive outcomes and/or selection.

At this point, readers might question whether engaging these models is worth the effort for outcomes such as distress or other factors commonly analyzed by behavioral scientists. Indeed, the time involved in undertaking these models for researchers not already familiar with them would be non-negligible. So then, is it worth the effort to undertake these models? In our opinion, it is. As more and more data sets become longitudinal, researchers, and manuscript reviewers, are becoming increasingly sensitive to issues of sample attrition. Consequently, even though accounting for attrition may yield little effect, readers will want to know that such analyses were undertaken to rule out the possibility of those effects. Likewise, even though we did

² Our estimates of incarceration rates were based on 1993 incarceration figures reported in the Snell (1995) report and the population size for African American men and women aged eighteen and older as reported in the 1990 census.

not find substantial differences in our own example, differences can exist. Without proper analytical techniques being applied, these differences could go undiscovered leading to erroneous findings and conclusions.

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